
Gestione della soglia di Rivelazione

Probabilità di falso allarme – soglia fissa

Sotto l'ipotesi H_0 ($a=0$) ho un Falso Allarme se

$$z = |\tilde{z}| = |d_f(t)| > T$$

Poiché $d_f(t)$ è Gaussiana a valor medio nullo e varianza σ_d^2 ($\sigma_d^2 = \sigma_n^2$ se disturbo = solo rumore termico):

$$p(\tilde{z} | H_0) = \frac{1}{\pi \sigma_d^2} \exp\left\{-\frac{1}{\sigma_d^2} |\tilde{z}|^2\right\} \quad \longrightarrow \quad p_z(z | H_0) = \frac{2z}{\sigma_d^2} e^{-\frac{z^2}{\sigma_d^2}}$$

$$P_{fa} = \text{Prob}\{z > T | H_0\} = \int_T^{\infty} p_z(z | H_0) dz = e^{-\frac{T^2}{\sigma_d^2}}$$

Scelta della soglia fissa

$$P_{fa} = e^{-\frac{T^2}{\sigma_d^2}}$$



$$\frac{T^2}{\sigma_d^2} = -\ln P_{fa}$$

* $T^2 \Big|_{dB} = \sigma_d^2 \Big|_{dB} + G_\infty$

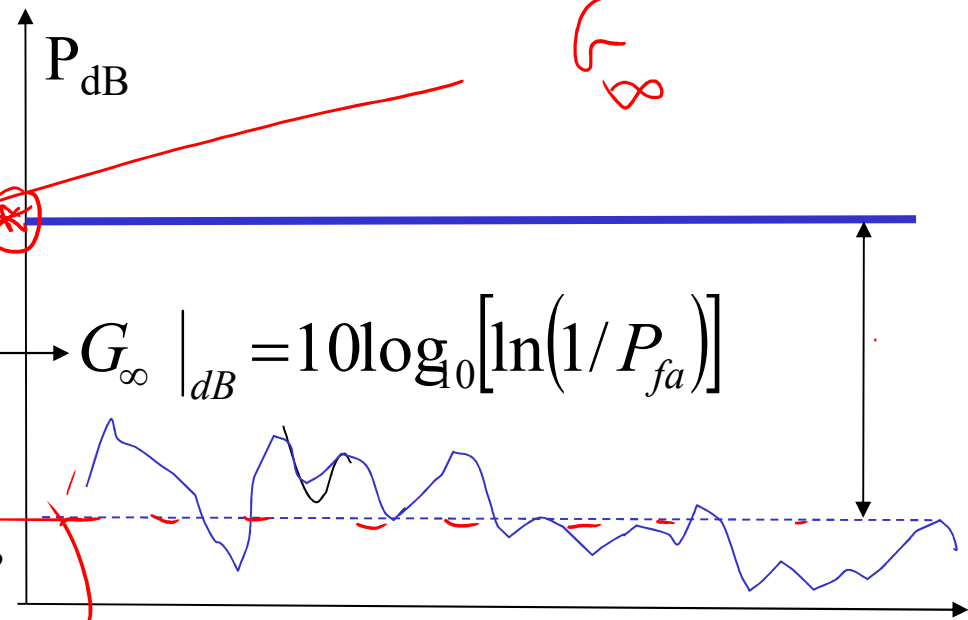
$$T = \sigma_d \cdot \sqrt{-\ln P_{fa}}$$



$$T^2 = \sigma_d^2 \cdot \ln(1/P_{fa})$$

$$G_\infty \Big|_{dB} = 10 \log_{10} [\ln(1/P_{fa})]$$

$$\sigma_d^2 \Big|_{dB}$$



Livello di potenza noto

Scelta della soglia – stima del disturbo

$$T^2 = \sigma_d^2 \cdot \ln(1/P_{fa})$$

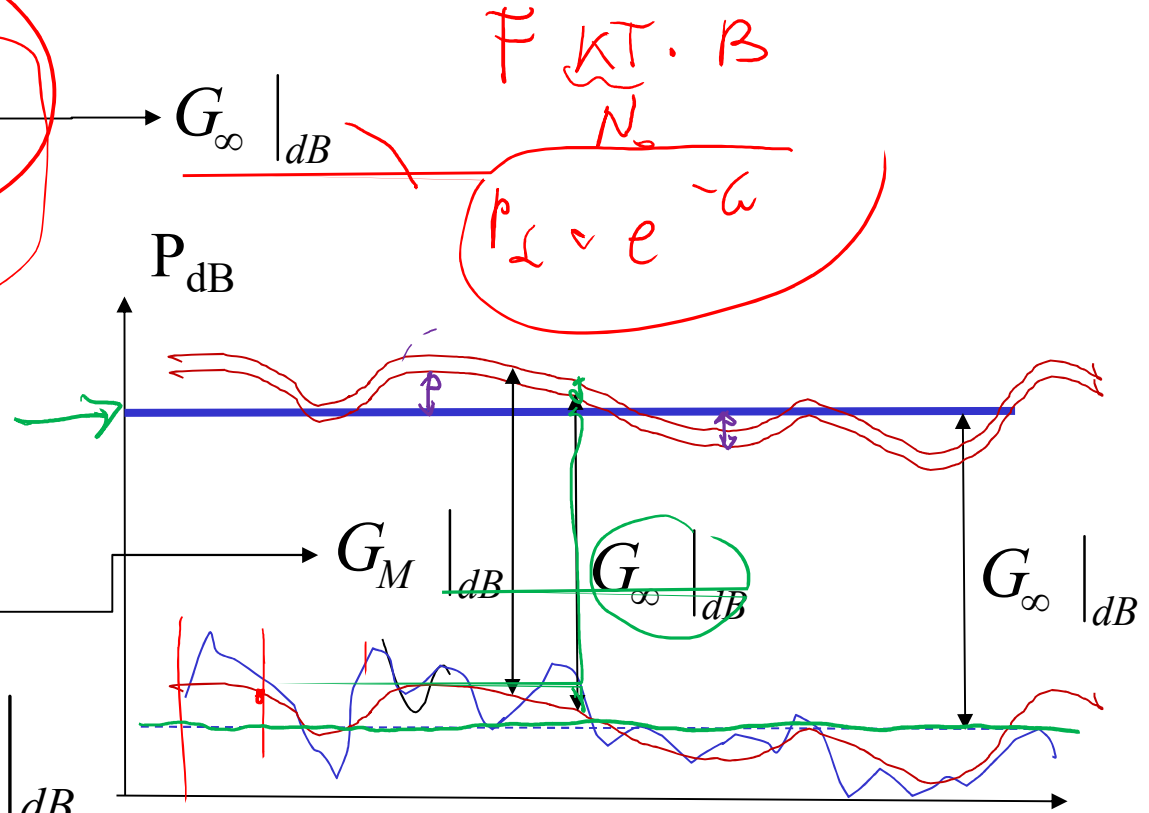
$$\sigma_d^2 |_{dB}$$

Livello di potenza noto

$$T^2 = \hat{\sigma}_d^2 \cdot G_M$$

$$\hat{\sigma}_d^2 |_{dB}$$

Livello di potenza stimato



$$P_d = e^{-G} \quad \ln P_d = -G$$

$$G = -\ln P_d$$

Incremento di P_{fa} per stima disturbo

$$\begin{aligned}
 P_{fa} &= \frac{1}{2} e^{-\frac{G(\sigma_d^2 + \Delta)}{\sigma_d^2}} + \frac{1}{2} e^{-\frac{G(\sigma_d^2 - \Delta)}{\sigma_d^2}} = \frac{1}{2} e^{-G \frac{\sigma_d^2 + \Delta}{\sigma_d^2}} + \frac{1}{2} e^{-G \frac{\sigma_d^2 - \Delta}{\sigma_d^2}} = \\
 &= e^{-G} \frac{1}{2} \left[e^{-\frac{G\Delta}{\sigma_d^2}} + e^{\frac{G\Delta}{\sigma_d^2}} \right] = e^{-G} \frac{1}{2} \left[1 + \frac{G\Delta}{\sigma_d^2} + \frac{1}{2} \left(\frac{G\Delta}{\sigma_d^2} \right)^2 + 1 + \frac{G\Delta}{\sigma_d^2} + \frac{1}{2} \left(\frac{G\Delta}{\sigma_d^2} \right)^2 \right] = \\
 &= e^{-G} \left[1 + \frac{1}{2} \left(\frac{G\Delta}{\sigma_d^2} \right)^2 \right] > 1
 \end{aligned}$$

con prob. $\frac{1}{2}$
 con valore noto

$$P_{fa} = e^{-\frac{\sigma_d^2 \cdot G}{\sigma_d^2}} = e^{-G}$$

$$\frac{1}{\sigma_d} = \sqrt{\sigma_d^2 + \Delta}$$

$$\frac{1}{\sigma_d} = \sqrt{\sigma_d^2 - \Delta}$$

P_{fa} per stima

$$P_{fa} = e^{-\frac{(\sigma_d^2 - \Delta)G}{\sigma_d^2}}$$

P_{fa} media?

$$T^2 = (\sigma_d^2 + \Delta) \cdot G$$

$P_{fa} = e^{-\frac{(\sigma_d^2 + \Delta)G}{\sigma_d^2}}$
 nel caso di stime per eccesso

$$T^2 = (\sigma_d^2 - \Delta) \cdot G$$

Autogate (I)

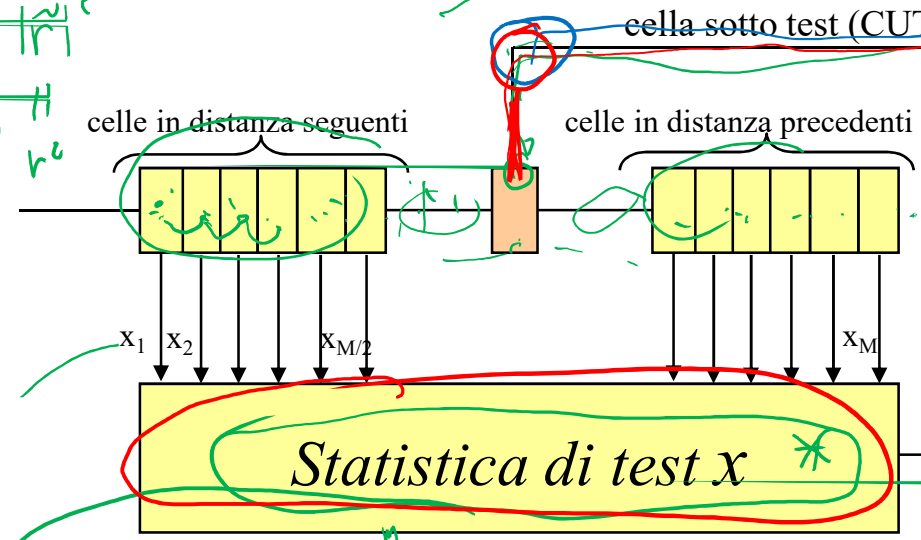
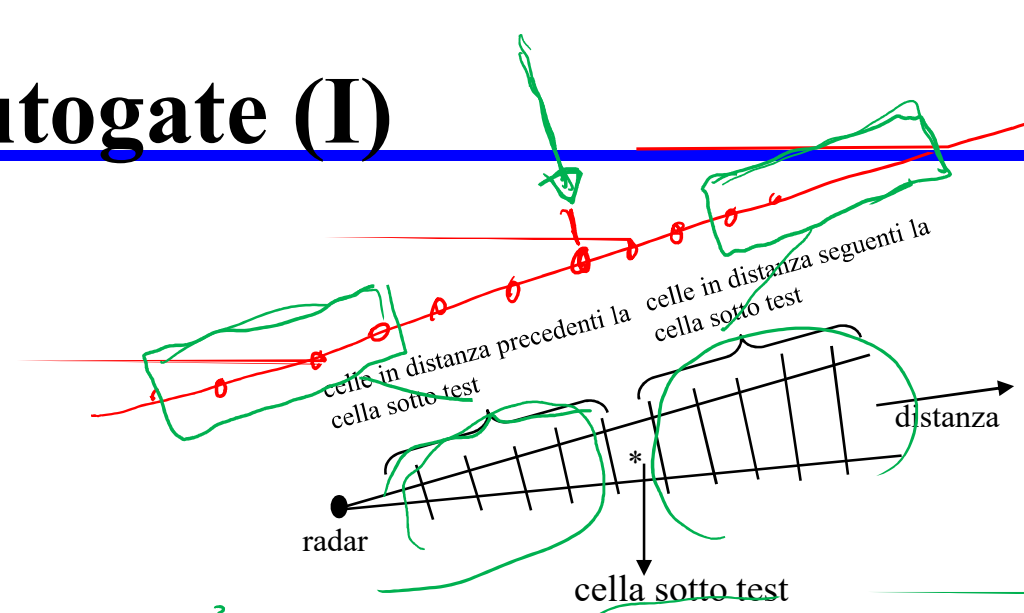
CFAR

(CA)-CFAR
CELL AVERAGE

rivelatore quadratico

$x_i = r_i^2$

$$\bar{x} = \frac{1}{M} \sum_{i=1}^M x_i$$



$r^2 \times G_m \times x$

$G \cdot x \leftrightarrow T$

G = Guadagno che determina la P_{fa}

PDF statistica di test (I) $x_i = r_i^2 = |\tilde{r}_i|^2$

Statistica di test

$$x = \frac{1}{M} \sum_{i=1}^M x_i = \frac{1}{M} x'$$

$$y = \alpha x$$

$$p_y(y) = \frac{1}{\alpha} p_x\left(\frac{y}{\alpha}\right)$$

$$x' = \sum_{i=1}^M x_i$$

- Ipotesi: disturbo gaussiano

$$p_y(y/H_0) = \frac{1}{\sigma^2} e^{-\frac{y}{\sigma^2}} \quad y \geq 0$$

$$p_{x_i}(x_i) = \frac{1}{\sigma^2} e^{-\frac{x_i}{\sigma^2}} \quad x_i \geq 0, i = 1 \dots M$$



M variabili aleatorie
indipendenti identicamente
distribuite

Per l'ipotesi di indipendenza statistica:
densità di probabilità = convoluzione delle singole
densità



$$p_{x'}(x') = p_{x_1}(x') * p_{x_2}(x') \dots * p_{x_M}(x')$$

Funzione caratteristica:

$$C_{x'}(u) = C_{x_1}(u) \cdot C_{x_2}(u) \cdot \dots \cdot C_{x_M}(u) = [C_{x_i}(u)]^M = \left[\frac{1/\sigma^2}{1/\sigma^2 - ju} \right]^M$$

antitrasformando

$$p_{x'}(x') = \frac{1}{(M-1)!} \frac{1}{\sigma^2} \left(\frac{x'}{\sigma^2} \right)^{M-1} e^{-\frac{x'}{\sigma^2}} \quad x' \geq 0$$

$$x = \frac{x'}{M} \Rightarrow p_x(x) = M p_{x'}(Mx)$$

$$C_{x_1}(\mu) = \int_{-\infty}^{+\infty} p_{x_1}(x') e^{-j2\pi\mu x'} dx'$$

$$C_{x_1}(\mu) \cdot C_{x_2}(\mu) \cdot C_{x_3}(\mu) \cdot \dots \cdot C_{x_M}(\mu) =$$

$$= \frac{1}{(1 + j2\pi\sigma^2\mu)^M}$$

$$C_{x_1}(\mu) \approx \int_0^{+\infty} \frac{1}{\sigma^2} e^{-\frac{x_1}{\sigma^2}} e^{-j2\pi\mu x_1} dx_1 = \frac{1}{\sigma^2} \frac{1}{1 + (j2\pi\mu\sigma^2)} = \frac{1}{1 + j2\pi\mu\sigma^2}$$

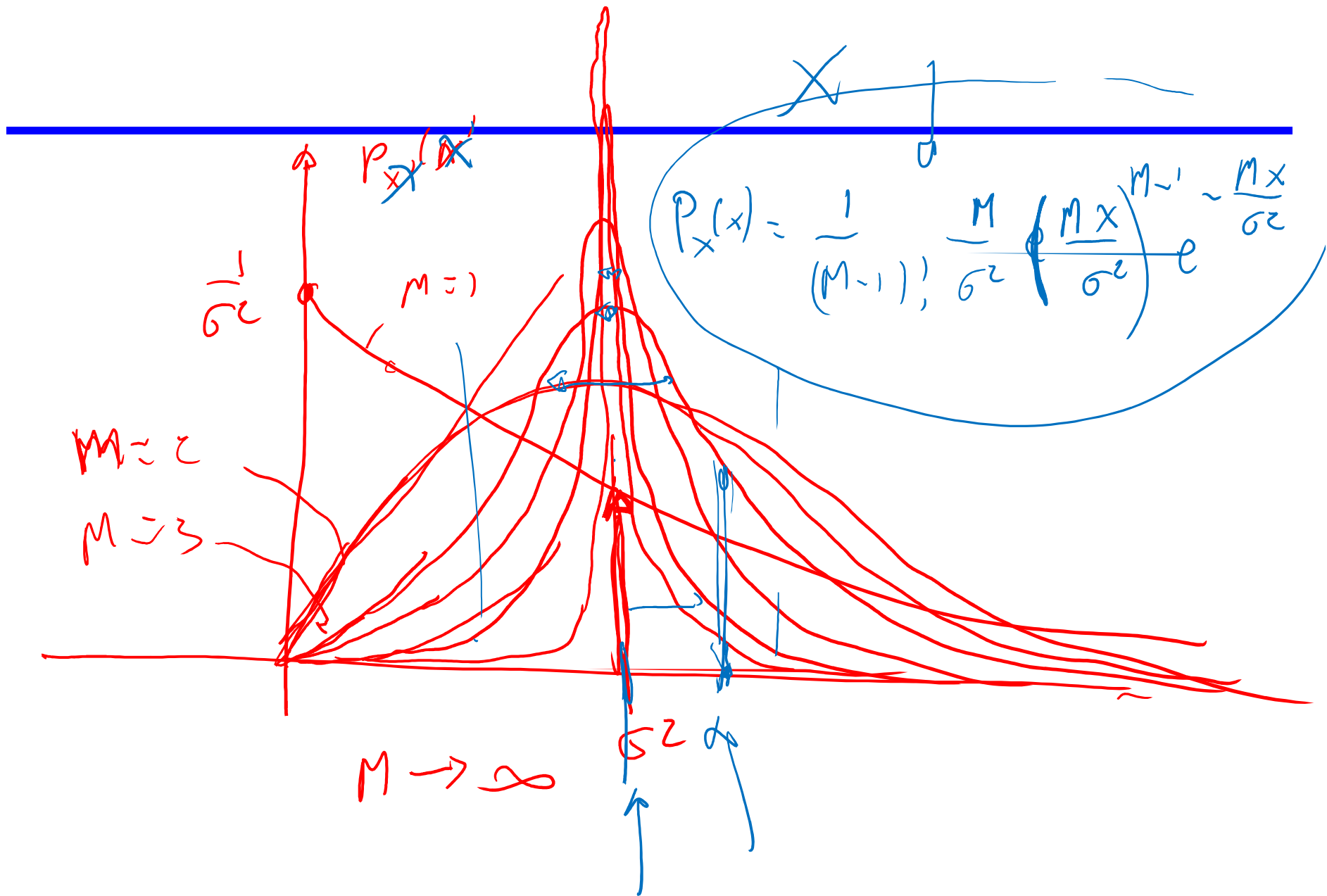
$$X' = \sum_{i=1}^M x_i = \sum_{i=1}^M |r_i|^2 = \sum_{i=1}^M (r_{iI}^2 + r_{iQ}^2)$$

Somma di $\sqrt{2M}$ ^{quadrati di} variabil. aleatorie gaussiane
 indip. con $\mu = 0$ e stessa varianza σ^2

$$P(x') = \frac{1}{(M-1)!} \frac{x'^{M-1}}{\sigma^2} e^{-\frac{x'}{\sigma^2}} \quad x' > 0$$

PDF GAMMA

$$M=1 \quad \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}}$$



DDP somma variabili esponenziali (I)

$$p_{x_i}(x_i) = \frac{1}{\sigma^2} e^{-\frac{x_i}{\sigma^2}} u(x_i)$$

$$p(x',2) = p_{x'}(x') * p_{x'}(x') = \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}} u(x') * \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}} u(x') = \int_0^\infty \frac{1}{\sigma^2} e^{-\frac{t}{\sigma^2}} u(t) \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} u(x'-t) dt =$$

$$= \left[\int_0^{x'} \frac{1}{\sigma^2} e^{-\frac{t}{\sigma^2}} \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} dt \right] u(x') = \left[\frac{1}{(\sigma^2)^2} e^{-\frac{x'}{\sigma^2}} \int_0^{x'} dt \right] u(x') = \frac{x'}{(\sigma^2)^2} e^{-\frac{x'}{\sigma^2}} u(x')$$

= 1 per $x' > t$
 $t < x'$

$$p(x',3) = p(x',2) * p_{x'}(x') = \frac{x'}{(\sigma^2)^2} e^{-\frac{x'}{\sigma^2}} u(x') * \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}} u(x') = \int_0^\infty \frac{t}{(\sigma^2)^2} e^{-\frac{t}{\sigma^2}} u(t) \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} u(x'-t) dt =$$

$$= \left[\int_0^{x'} \frac{t}{(\sigma^2)^2} e^{-\frac{t}{\sigma^2}} \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} dt \right] u(x') = \left[\frac{1}{(\sigma^2)^3} e^{-\frac{x'}{\sigma^2}} \int_0^{x'} t dt \right] u(x') = \frac{x'^2}{2(\sigma^2)^3} e^{-\frac{x'}{\sigma^2}} u(x')$$

$$p(x',4) = p(x',3) * p_{x'}(x') = \frac{x'^2}{2(\sigma^2)^3} e^{-\frac{x'}{\sigma^2}} u(x') * \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}} u(x') = \int_0^\infty \frac{t^2}{2(\sigma^2)^3} e^{-\frac{t}{\sigma^2}} u(t) \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} u(x'-t) dt =$$

$$= \left[\int_0^{x'} \frac{t^2}{2(\sigma^2)^3} e^{-\frac{t}{\sigma^2}} \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} dt \right] u(x') = \left[\frac{1}{2(\sigma^2)^4} e^{-\frac{x'}{\sigma^2}} \int_0^{x'} t^2 dt \right] u(x') = \frac{x'^3}{3 \cdot 2(\sigma^2)^4} e^{-\frac{x'}{\sigma^2}} u(x')$$

DDP somma variabili esponenziali (II)

$$\begin{aligned}
 p(x',5) &= p(x',4) * p_{x'}(x') = \frac{x'^3}{3 \cdot 2(\sigma^2)^4} e^{-\frac{x'}{\sigma^2}} u(x') * \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}} u(x') = \int_0^\infty \frac{t^3}{3 \cdot 2(\sigma^2)^4} e^{-\frac{t}{\sigma^2}} u(t) \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} u(x'-t) dt = \\
 &= \left[\int_0^{x'} \frac{t^3}{3 \cdot 2(\sigma^2)^4} e^{-\frac{t}{\sigma^2}} \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} dt \right] u(x') = \left[\frac{1}{3 \cdot 2(\sigma^2)^5} e^{-\frac{x'}{\sigma^2}} \int_0^{x'} t^3 dt \right] u(x') = \frac{x'^4}{4 \cdot 3 \cdot 2(\sigma^2)^5} e^{-\frac{x'}{\sigma^2}} u(x')
 \end{aligned}$$

$$p(x',M) = \frac{1}{(M-1)!} \frac{x'^{M-1}}{(\sigma^2)^M} e^{-\frac{x'}{\sigma^2}} u(x')$$

Gamma

$$\begin{aligned}
 p(x',M+1) &= p(x',M) * p_{x'}(x') = \frac{1}{(M-1)!} \frac{x'^{M-1}}{(\sigma^2)^M} e^{-\frac{x'}{\sigma^2}} u(x') * \frac{1}{\sigma^2} e^{-\frac{x'}{\sigma^2}} u(x') = \\
 &= \int_0^\infty \frac{1}{(M-1)!} \frac{t^{M-1}}{(\sigma^2)^M} e^{-\frac{t}{\sigma^2}} u(t) \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} u(x'-t) dt = \left[\int_0^{x'} \frac{1}{(M-1)!} \frac{t^{M-1}}{(\sigma^2)^M} e^{-\frac{t}{\sigma^2}} \frac{1}{\sigma^2} e^{-\frac{x'-t}{\sigma^2}} dt \right] u(x') = \\
 &= \left[\frac{1}{(M-1)!} \frac{1}{(\sigma^2)^{M+1}} e^{-\frac{x'}{\sigma^2}} \int_0^{x'} t^{M-1} dt \right] u(x') = \frac{1}{M!} \frac{x'^M}{(\sigma^2)^{M+1}} e^{-\frac{x'}{\sigma^2}} u(x')
 \end{aligned}$$

DDP somma variabili esponenziali (III)

$$\int_0^{\infty} p(x', M) dx' = \int_0^{\infty} \frac{1}{(M-1)!} \frac{x'^{M-1}}{(\sigma^2)^M} e^{-\frac{x'}{\sigma^2}} u(x') dx' = 1$$

$$m_{x'}^k = \int_0^{\infty} x'^k p(x', M) dx' = \int_0^{\infty} x'^k \frac{1}{(M-1)!} \frac{x'^{M-1}}{(\sigma^2)^M} e^{-\frac{x'}{\sigma^2}} u(x') dx' =$$

$$= \frac{(M+k-1)!}{(M-1)!} \frac{1}{(\sigma^2)^k} \int_0^{\infty} \frac{1}{(M+k-1)!} \frac{x'^{M+k-1}}{(\sigma^2)^{M+k}} e^{-\frac{x'}{\sigma^2}} u(x') dx' = 1$$

$$= \frac{(M+k-1)!}{(M-1)!} (\sigma^2)^k = (M+k-1) \cdot (M+k-2) \cdot \dots \cdot M \cdot (\sigma^2)^k$$

$$m_{x'}^1 = M \cdot \sigma^2$$

$$m_{x'}^2 = (M+1) \cdot M \cdot (\sigma^2)^2$$

$$\text{var} = m_{x'}^2 - (m_{x'}^1)^2 = (M+1) \cdot M \cdot (\sigma^2)^2 - M^2 \cdot (\sigma^2)^2 = M \cdot (\sigma^2)^2$$

$$C = \frac{\text{var}}{(m_{x'}^1)^2} = \frac{m_{x'}^2 - (m_{x'}^1)^2}{(m_{x'}^1)^2} = \frac{M \cdot (\sigma^2)^2}{M^2 \cdot (\sigma^2)^2} = \frac{1}{M}$$

DDP media variabili esponenziali (I)

$$x = \frac{1}{M} x'$$

$$p(x, M) = M \cdot p(Mx, M)$$

$$p(x, M) = \frac{M}{(M-1)! (\sigma^2)^M} (Mx)^{M-1} e^{-\frac{Mx}{\sigma^2}} u(x) = \frac{1}{(M-1)! \left(\frac{\sigma^2}{M}\right)^M} x^{M-1} e^{-\frac{Mx}{\sigma^2}} u(x)$$

$$m_x^k = E\left\{\left(\frac{1}{M} x'\right)^k\right\} = \frac{1}{M^k} m_{x'}^k = \frac{(M+k-1)! (\sigma^2)^k}{M^k (M-1)!} = \frac{(M+k-1) \cdot (M+k-2) \cdot \dots \cdot (M+1) \cdot M \cdot (\sigma^2)^k}{M^k} =$$

$$= \left(1 + \frac{k-1}{M}\right) \cdot \left(1 + \frac{k-2}{M}\right) \cdot \dots \cdot \left(1 + \frac{1}{M}\right) \cdot 1 \cdot (\sigma^2)^k$$

$$m_x^1 = \sigma^2$$

$$m_x^2 = \left(1 + \frac{1}{M}\right) \cdot (\sigma^2)^2$$

$$\text{var} = m_x^2 - (m_x^1)^2 = \left(1 + \frac{1}{M}\right) \cdot (\sigma^2)^2 - (\sigma^2)^2 = \frac{1}{M} (\sigma^2)^2$$

$$C = \frac{\text{var}}{(m_x^1)^2} = \frac{m_x^2 - (m_x^1)^2}{(m_x^1)^2} = \frac{1}{M} \frac{(\sigma^2)^2}{(\sigma^2)^2} = \frac{1}{M}$$

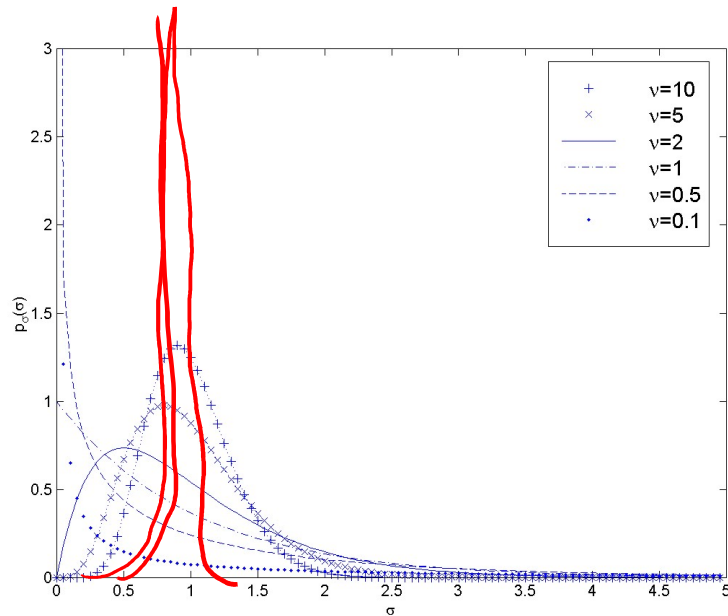
La DDP gamma

Simple solution is given by the gamma PDF:

$$p(t) = \left(\frac{\nu}{\sigma^2}\right)^\nu \frac{t^{\nu-1}}{\Gamma(\nu)} \exp\left[-\frac{\nu}{\sigma^2} t\right]$$

where n is an order parameter, with moments given by

$$\langle t^n \rangle = \left(\frac{\sigma^2}{\nu}\right)^n \frac{\Gamma(n+\nu)}{\Gamma(\nu)}$$



The square of the contrast or CV, i.e. the normalised variance, is

$$\text{var } t / (\sigma^2)^2 = 1/\nu$$

Autogate (II)

$$* P_d = \int_T^{\infty} p(r/H_0) dy$$

Probabilità di falso allarme:

$$P_{fa} = \int_0^{\infty} \left[\int_{G \cdot x}^{\infty} p_y(y/H_0) dy \right] p_x(x) dx$$

$$P_d(x) = \int_{G \cdot x}^{\infty} p(y/H_0) dy$$

P_d/x

DDP della cella sotto test

DDP della statistica di test

Probabilità di rivelazione:

$$P_d = \int P_d(x) p_x(x) dx$$

$$\int p(z/x) \cdot p_x(x) dx$$

$$P_d = \int_0^{\infty} \left[\int_{G \cdot x}^{\infty} p_y(y/H_1) dy \right] p_x(x) dx$$

CA-CFAR (II)

$$\left[-e^{-\frac{y}{\sigma^2}} \right]_{Gx}^{\infty} = e^{-\frac{Gx}{\sigma^2}} \quad e^{-\frac{I}{\sigma^2}}$$

$$P_{fa} = \int_0^{\infty} \left[\int_{Gx}^{\infty} \frac{1}{\sigma^2} e^{-\frac{y}{\sigma^2}} dy \right] p_x(x) dx = \int_0^{\infty} e^{-\frac{Gx}{\sigma^2}} p_x(x) dx = \int_0^{\infty} e^{-\frac{Gx}{\sigma^2}} \frac{M-1}{(M-1)! \sigma^2} \left(\frac{Mx}{\sigma^2} \right)^{M-1} e^{-\frac{Mx}{\sigma^2}} dx =$$

$$= \int_0^{\infty} \frac{M-1}{(M-1)! \sigma^2} \left(\frac{Mx}{\sigma^2} \right)^{M-1} e^{-\frac{x}{\sigma^2}(M+G)} dx = \left(\frac{M-1}{M+G} \right)^{M-1} \int_0^{\infty} \frac{1}{(M-1)! \sigma^2} \left(\frac{z}{\sigma^2} \right)^{M-1} e^{-\frac{z}{\sigma^2}} dz = \left(1 + \frac{G}{M} \right)^{-M}$$

$$\frac{M-1}{M+G} \left(\frac{M+G}{M} \right)^{M-1} \int_0^{\infty} \frac{1}{(M-1)! \sigma^2} \left(\frac{z}{\sigma^2} \right)^{M-1} e^{-\frac{z}{\sigma^2}} dz$$

Integrale di una densità di probabilità $z = (M+G)x$

Fissato il tasso di falsi allarmi desiderato e fissato il valore di M resta fissato il valore da dare al guadagno G.

$$\frac{M-1}{M+G} \left(\frac{M+G}{M} \right)^{M-1} = \left(\frac{M-1}{M+G} \right)^{M-1} \left(\frac{M+G}{M} \right)^{M-1} = \left(\frac{M-1}{M} \right)^{M-1} \left(\frac{M+G}{M} \right)^M = \left(\frac{1}{1 + \frac{G}{M}} \right)^M = \frac{1}{\left(1 + \frac{G}{M} \right)^M}$$

Probabilità di falso allarme dipende solo da G e da M ed è indipendente dal livello del disturbo \Rightarrow **CFAR: Constant False Alarm Rate** indipendente da variazioni spaziali o temporali del livello del disturbo (la soglia si adatta a variazioni spaziali e/o temporali di σ^2).

$$P_{\alpha}(M) = \left(1 + \frac{G}{M}\right)^{-M}$$

$$P_{\alpha}^{-1/M} = 1 + \frac{G}{M}$$

$$G_M = M \cdot \left[P_{\alpha}^{-1/M} - 1 \right] \leftarrow$$

Al crescere di M G_M decresce

$$\lim_{M \rightarrow \infty} P_{\alpha}(M) = \lim_{M \rightarrow \infty} \left(1 + \frac{G}{M}\right)^{-M} = e^{-G}$$

$$P_{\alpha} = e^{-\frac{G \cdot \cancel{\sigma^2}}{\sigma^2}}$$

CA-CFAR (III)

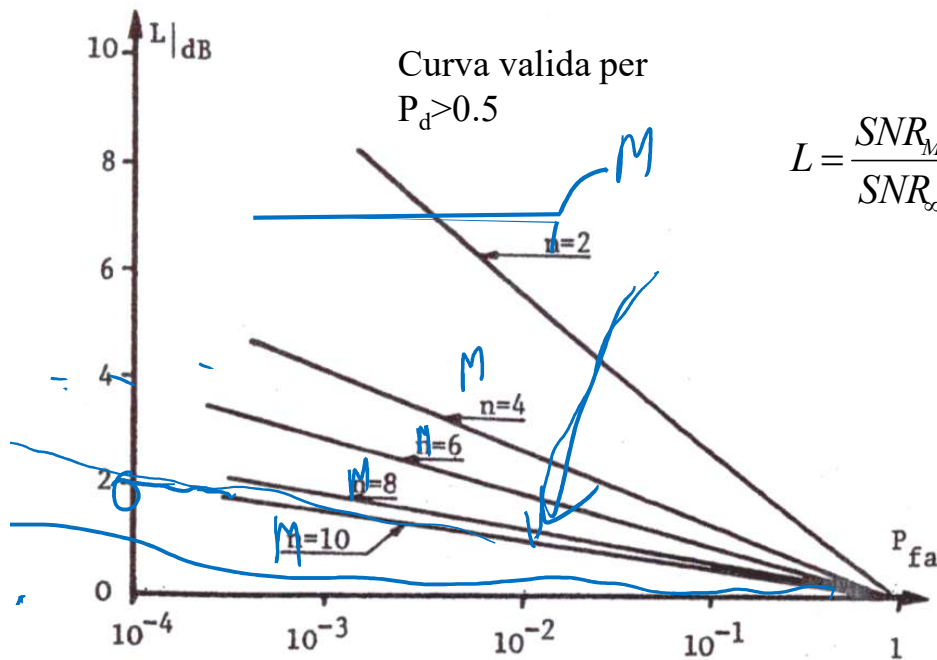
- Probabilità di rivelazione:

$$P_d = \int_0^\infty \left[\int_{G_x}^\infty p_y(y/H_1) dy \right] p_x(x) dx$$

$P_d = 10^{-6}$
 $P_d = 0,9$

$\frac{S}{N} = 13 \text{ dB}$

- soglia fluttuante → perdite in sensibilità (tutto va come se ci fosse una soglia fissa più un rumore additivo dovuto alla fluttuazione della stima)



$$L = \frac{SNR_M}{SNR_\infty}$$

Perdite in sensibilità: rapporto tra SNR richiesto in presenza di soglia stimata adattivamente da M campioni (SNR_M) e SNR richiesto in presenza di soglia fissa (SNR_∞).

Fissato il livello di falso allarme le perdite sono tanto minori quanto maggiore è M (M=∞ equivale a disturbo noto).

Loss 4 dB → 17 dB