Full-Length Research Report

Commitment and Dyadic Coping in Long-Term Relationships

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Abstract. *Background:* Previous research focused on relationship commitment as the outcome of high satisfaction, poor alternatives, and high investments. We propose that commitment is a prerequisite in highly satisfied couples, fostering relationship maintenance behavior such as positive dyadic coping. *Method:* Structural equation models identified the relationship between commitment, relationship satisfaction, and dyadic coping with data from 201 heterosexual couples with an average relationship length of 34 years. *Results:* The common fate model confirmed that relationship satisfaction mediated the effects between commitment and dyadic coping on a latent dyadic level. Additional analyses revealed that women's satisfaction was mainly responsible for mediating effects between both partners' commitment and dyadic coping. *Conclusions:* Findings support the essentiality of commitment for couples' maintenance strategies and for consistency in long-term relationships.

Keywords: commitment, dyadic coping, relationship satisfaction, dyadic analysis, couples

Introduction

Commitment is known to foster a wide range of relationship-building and relationship-maintaining behavior, and is decisive to relationship continuity because it stands for the willingness of the intimate partners to actively work on their relationship (Johnson, 1991; Rusbult & Buunk, 1993). Here, we explore the relationship between dyadic partnership variables in long-term relationships. To this end we investigated a direction of the association between commitment and prorelationship behavior Ogolsky (2009) called the "motivational model." Whereas relationship commitment is preponderantly considered a criterion variable (e.g., Canary & Stafford, 1992; Floyd & Wasner, 1994; Rusbult, 1983), the motivational model postulates that commitment in long-term relationships can function as a predictor and lead to increased levels of prorelationship behavior. We suggest that commitment predicts prorelationship behavior in the form of common dyadic coping, and that this association is mediated by relationship satisfaction. As high commitment was found to trigger a wide range of relationship-favorable behaviors in individual (Rusbult, Olsen, Davis, & Hannon, 2001; Rusbult, Verette, Whitney, Slovik, & Lipkus, 1991; Van Lange, Agnew, Harinck, & Steemers, 1997; Van Lange, Rusbult et al., 1997) as well as in dyadic analyses (Weigel & Ballard-Reisch, 2008), we believe that relationship commitment can function as a starting point in a relationship, and that many behavior patterns within an intimate partnership are influenced by it.

Relationship Commitment

In an intimate relationship, relational commitment arises from both previous and existing experiences with partnership dependence (Rusbult, 1980). It also functions as a trendsetter for further directions in the partners' interdependence behavior (Tran & Simpson, 2009). Interdependence theory (Thibaut & Kelley, 1959) suggests that partners are dependent on one another reciprocally inasmuch as the partners fulfill their respective needs. It is this mutual need for "instrumental support, affection, sexual fulfillment, and emotional closeness" (Rusbult & Buunk, 1993, p. 179) which intensifies the dependence of intimate partners and nurtures their relationship satisfaction. However, the authors make a distinction between levels of satisfaction and levels of dependence: The former stands for high levels of appreciation of the relationship conditions and for valuing partners' positive and need-fulfilling behavior; the latter reflects a relationship that is based primarily on need fulfillment. The authors expand the interdependence theory by introducing commitment and formulating the investment model of commitment. This model suggests that "dependence is subjectively represented by feelings of commitment" (Rusbult, Drigotas & Verette, 1994, p. 119). Thus, whenever couples reach the level beyond individual need fulfillment, they may feel a satisfaction that rest not only on their own rewards, but also on the rewards and need fulfillment of their partner.

Commitment includes the long-term perspective to enter and sustain an intimate relationship (Frank & Brandstätter, 2002; Sternberg, 1986; Wieselquist, Rusbult, Foster, & Agnew, 1999). Thus, because of an ongoing relationship commitment that is maintained by the dependence on and the need for relationship satisfaction, intimate partners have a particularly high interest in applying a wide range of prorelationship behaviors. Rusbult et al. (1994) define commitment as a "macromotive" (p. 123) for partnership, which then triggers prorelationship behavior. For instance, highly committed individuals are less inclined to retaliate against their intimate partners' provocative behavior and more prepared to forgive negative acts than less committed individuals (Finkel, Rusbult, Kumashiro, & Hannon, 2002). According to the authors, strongly committed individuals have a stronger sense of "we-ness" and consciously pursue the objective of a long-term relationship. They are therefore inclined to "develop patterns of reciprocal pro-relationship behavior" (p. 96).

Accordingly, the higher the commitment, the more stable and satisfactory the relationship, and the higher the partners' willingness to intensify their relational maintenance efforts for mutual adjustment (Rusbult, 1983; Schneewind & Wunderer, 2003). These interactions and interdependencies result in a reinforcing loop in which commitment can function as both predictor and criterion (see Figure 1).

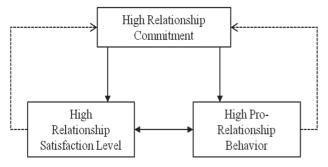


Figure 1. The reinforcing loop model of relationship commitment, based on Rusbult's investment model of commitment (1983).

Dyadic Coping: A Form of Prorelationship Behavior

Relational maintenance stands for recurring dynamic affinity-enhancing activities that intimate partners undertake to keep their relationship on a satisfactory level (Bell, Daly, & Gonzalez, 1987) and is inherently connected to intimate relational constructs such as partnership satisfaction, partnership stability, and commitment (Canary & Stafford, 1994). According to Bell et al. (1987), typical prorelationship processes include – among other – altruism, listening, optimism, sensitivity, supportiveness, and verbal affection. These constructs account for some of the major characteristics of dyadic coping (for more details on dyadic coping, see Bodenmann, 1995, 2005). Positive supportive dyadic coping, for example, is defined as a partner's supportive reactions toward the other partner's stress signals, such as empathic understanding, showing solidarity, or being encouraging (Bodenmann, 2005). Supportive dyadic coping thus unites precisely those aspects of prorelationship tendencies that define one partner's efforts to appease the other partner's stress by assisting her or him in their coping efforts through providing emotional and problem-focused support (Meuwly et al., 2012). Negative dyadic coping, on the other hand, can be seen as corresponding to what Rusbult and Zembrodt (1983) call neglect behavior (i.e., not valuing or appreciating the partner, criticizing the partner, or ignoring the partner in times of dissatisfaction). In our study, we use common dyadic coping, which is a subform of positive dyadic coping (Bodenmann, Charvoz, Widmer, & Bradbury, 2004). It describes a common process that involves the reciprocal behavior of both partners, including joint appraisals, feedback, and joint problem management (Lazarus, 1999).

Research has provided evidence for a positive connection between levels of commitment and prorelationship behavior (Reis & Collins, 2000; Van Lange, Agnew et al., 1997; Van Lange, Rusbult et al., 1997), on the one hand, and dyadic coping and relationship satisfaction on the other (Bodenmann & Cina, 2006; Bodenmann, Pihet, & Kayser, 2006; Bodenmann, Meuwly, & Kayser, 2011; Herzberg, 2013; Landis, Peter-Wight, Martin, & Bodenmann, 2013; Papp & Witt, 2010). In line with these findings we speculate that commitment brings forth the willingness to positively cope dyadically. In the present study, our focus lies on conjoint or common dyadic coping, where intimate partners manage a situation by combining their coping efforts. It is thus a conjoint effort made by both partners to be responsive to each other's stress reactions and to resolve upcoming problems as a couple.

The specification of this variable is essential as it takes into account the nonindependence of couples' data and, hence, fulfills an important prerequisite for our methodological approach, which models intimate partners' conjoint coping as a common fate factor (Ledermann, Bodenmann, Rudaz, & Bradbury, 2010; Ledermann & Kenny, 2011).

The Current Study

In this project, we demonstrate that intimate couples who are highly committed to their relationship and who show particularly high levels of relationship satisfaction also show high levels of common dyadic coping (Hypothesis 1). This relationship between the variables might seem self-evident, but though high and stable levels of commitment can be considered a prerequisite to a long-term relationship, commitment does not necessarily or directly lead to high levels of relationship satisfaction. In fact, high levels of commitment to a partner who frequently requires individual and dyadic coping skills may cause a lower relationship satisfaction. In that case, commitment might still be positively related to positive coping (resulting in one partner reporting higher levels of supportive positive coping), and at the same time be negatively related to their relationship satisfaction. In our study, however, we pay special attention to conjoint coping efforts, reported by both partners, expressing the couples' willingness to invest in their relationship.

Commitment, relationship satisfaction, and dyadic coping are understood as between-dyad variables that reflect a common-fate construct and are relationship-referential (see Ledermann & Kenny, 2011), as both members of the dyad have to assess the variables and their assessment is not independent. Following the demand by Thompson-Hayes and Webb (2004) to treat commitment as a dyadic variable, we used the common fate model (CFM; Ledermann & Macho, 2009) to calculate our hypothesized mediation model, which stated that relationship satisfaction would mediate the association between relational commitment and common dyadic coping on a dyadic level (*Hypothesis* 2). The CFM is especially suited for this mediation analysis because it reflects the nature of between-dyad variables for which the dyad is the unit of analysis. In this case, we were interested in assessing a variable from both partners while taking into account that their scores are not perfectly congruent.

In addition, as dyadic coping and commitment can also be understood as personal variables assessing individual behavior, we were particularly interested in understanding the mediation process on the level of dyad members, and we used the actor-partner-interdependence mediation model (API-MeM; Ledermann & Bodenmann, 2006), an extension of the APIM, for further analyses. The APIMeM consists of two exogenous variables and two endogenous variables, which are linked by the two mediator variables. Its purpose is to show that significant associations can exist between exogenous variables and endogenous variables, between exogenous variables and potential mediators, and between the mediators and the endogenous variables (Ledermann & Bodenmann, 2006). Because of prior evidence for women's higher emotional involvement in their relationship (Anderson, Keltner, & John, 2003), and because of women's higher probability to transfer personal dysphoria onto marital stress (Davila, Bradbury, Cohan, & Tochluk, 1997), as well as because of their increased vigilance to fluctuations in their relationship (Nolen-Hoeksema & Jackson, 2001), we expected women's relationship satisfaction to also influence men's variables in our setting. In other words, we hypothesized that women's relationship satisfaction would be a stronger mediator between both partners' commitment and dyadic coping than men's (*Hypothesis 3*).

In order to analyze whether age and relationship length moderated the mediating effects in the CMF model and the APIMeM, we additionally ran multigroup analyses with age cohort or relationship length (median split) as grouping variable. Since we did not have any specific hypothesis regarding differences with respect to age or relationship length, we tested a completely restricted multigroup model first with all model parameters (except for means and intercepts) to be equal across groups. If this model would fit the data, there would be no structural differences between the different groups. If this model would not fit the data, restrictions could be cleared from the model in the following order: residuals and path coefficients.

Methods

Participants

Data for this study were extracted from a larger study in which a total of 368 intimate couples were recruited by means of newspaper articles and advertisements. Participation criterion required couples to have been sharing an intimate relationship for at least 1 year at the time of survey. Couples were recruited in three different age cohorts: (1) ranging from 20 to 35 years, (2) ranging from 40–55 years, and (3) ranging from 65–80 years. In the present study, we had a special interest in examining couples with longer relationship durations. As age turned out to be highly correlated with relationship duration, we concentrated on cohorts 2 and 3 and on couples with a minimal relationship duration of 10 years. Our dataset initially contained a total of 246 dyads for cohorts 2 and 3. After exclusion because of missing data and relationship duration condition, our final sample included 201 dyads (92 in the younger age group and 109 in the older age group) whose average relationship duration was M = 33.74 years (SD = 14.17), ranging from 10 to 60 years, and representing a typical sample representing a typical sample of nonclinical relationships (Bodenmann et al., 2011; Sprenkle & Olson, 1978). The mean age for women was 57.97 (SD = 12.56), men's average age was 60.38 (SD = 12.30).

Measures

Cognitive and Emotional Commitment

As Rusbult and Buunk (1993) state, commitment includes "both cognitive and emotional components" (p. 180).

Therefore the use of the ComSec (Bodenmann & Kessler, 2011, unpublished manuscript) with its subscales *emotional commitment* and *cognitive commitment* seemed highly appropriate to function as indicators of the latent construct *partnership commitment* (including both aspects, emotional and cognitive commitment). Couples rated two cognitive commitment items ("It is my goal to grow old together with my partner" and "It is my goal to make our partnership last forever") and two emotional commitment items (e.g., "It is my goal to get fully involved emotionally with my partner") on a 7-point Likert scale ranging from $1 = does \ not \ apply \ at \ all \ to \ 7 = is \ absolutely \ true)$. Cronbach's α was .80 for women's commitment and .80 for men's commitment for the total score.

Relationship Satisfaction

We used the German version (Sander & Böcker, 1993) of the Relationships Assessment Scale (Hendrick, 1988). Couples rated their relationship satisfaction on seven items in a 5-point Likert format, ranging from 1 = not at all to 5 = completely, example items being "How well does your partner meet your needs?" and "To what extent has your relationship met your original expectations?"; negative items 4 and 7 were reverse-score recoded. Cronbach's α was .89 for women's and .90 for men's scale.

Dyadic Coping

The Dyadic Coping Inventory (DCI, Bodenmann, 2008) is a self-report questionnaire based upon the systemic-transactional stress concept by Bodenmann (1997). It comprises items related to (1) the expression of stress signals by one partner and (2) the other partner's corresponding responsive reactions, namely, as defined by supportive, negative, and common dyadic coping, each of the forms being subdivided into problem- and emotion-focused support. It consists of 37 items that can be answered from 1 = hardly everto 5 = very often. Both partners answer the questionnaire individually – male and female questionnaires are identical in items but gender-adapted. The questionnaire consists of the following scales: (1) own stress communication (emotional, problem-oriented, four items), (2) own supportive dyadic coping (emotional, problem-oriented, delegated, seven items), (3) own negative dyadic coping (hostile, ambivalent, withdrawal, four items), (4) own evaluation of conjoint dyadic coping (satisfaction with dyadic coping, efficiency of dyadic coping, five items), (5) partner's stress communication (emotional, problem-oriented, four items,), (6) partner's supportive dyadic coping (emotional, problem-oriented, delegated, seven items), (7) partner's negative dyadic coping (hostile, ambivalent, withdrawal, four items), and finally two items that evaluate the satisfaction with and the efficiency of the partner's coping support. Sum

scores can be calculated for all above-mentioned scales. Additionally, summarizing the scores for supportive dyadic coping, stress communication, and negative dyadic coping (reversed polarity) yields the total score for dyadic coping for the 35 items ranges from 35 to 175 points.

For our study, we concentrated on three of the five items assessing the conjoint coping efforts, leaving out the two items with sexual connotations and concentrating on those assessing mutual comforting and exchange of relevant information on the stress event (see Bodenmann et al., 2004). Sample items were "We try to handle the problem together and to come up with specific solutions" (item 31), "We engage in a serious discussion about the problem and think through what has to be done" (item 32), "We help each other to see the problem in a new light" (item 33). We opted for these items as we had also decided to not include the items on sexual commitment of the ComSec, leaving out the sexual components in our study. This decision was based on factor analysis conducted for both variables; the rotated component matrix clearly distinguished the sexual items from the other items. Cronbach's α was .87 for the three items of women's assessment of common dyadic coping, and .90 for men's.

Data Analyses

In addition to correlational analyses, which we performed with IBM SPSS Statistics (Version 20), we used Mplus 7.2 (Muthén & Muthén, 1998-2012) to estimate the common fate model (CFM, Figure 2; Ledermann & Kenny, 2011) as well as the actor-partner-interdependence-model of mediation (APIMeM, Ledermann, Macho, & Kenny, 2011) as structural equation models (SEM). The CFM analysis uses scores from both dyad members and integrates them into the reflective construct. In this model, latent dyadic variables are defined by the indicator scores of both husbands and wives on the same variable (in our model commitment, dyadic coping, and relationship satisfaction). The CFM states that both members of the dyad are influenced by a shared latent construct that affects both, which makes it appropriate in the case of dyad members assessing a variable that expresses their common behavior. In this study, the items assessing common dyadic coping fully meet this criterion, as they render the spouses' assessments of their joint efforts to cope with stress affecting them as a couple. Example items are "We seriously consider the problem and analyse what needs to be done" and "We help each other to reconsider the problem in a new light." The same prerequisite criterion is met by relationship satisfaction; the variable is per se nonindependent of micro- and macrocontextual processes affecting both partners (see Bradbury, Fincham, & Beach, 2000, for an overview on determinants of marital satisfaction), and for relationship commitment. According to Ledermann and Macho (2009), three basic assumptions must be met in order to choose the common

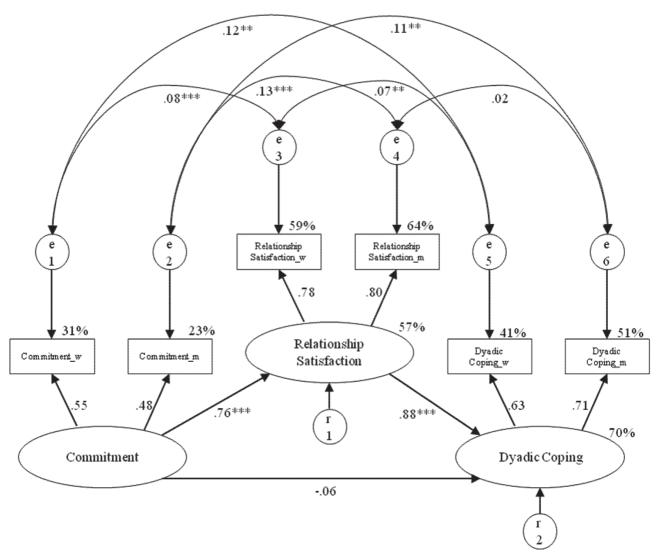


Figure 2. The common fate mediation model with standardized coefficients testing the association between commitment, relationship satisfaction and dyadic coping. Note: w = women; m = men. Percentages denote explained variances. *p < .05, **p < .01, ***p < .001.

fate mediation model for data analysis: First, the dyadmembers' data are influenced by a latent reflective variable. Second, there is evidence for mediation on the dyadic level. Third, the scores of husbands and wives are true indicators of the latent constructs. In our study these assumptions can be made: Commitment and relationship satisfaction are variables that qualify the intimate relationship and are common dyadic variables influencing both partners, respectively. Furthermore, dyadic coping – specifically the items assessing common dyadic coping – is a prime example of a dyadic construct which affects both partners and which renders the dyad-members assessment of their dyadic behavior (see section on measurements).

While the common fate model integrates the scores of both partners, displaying the associations between the latent dyadic constructs, the APIMeM takes into account the nonindependent nature of dyadic data and uncovers interpersonal as well as intrapersonal associations between variables in distinguishable dyads.

Results

Table 1 lists the means and standard deviations of the study variables. Women and men differed significantly in their appraisal of common dyadic coping efforts as well as in relationship satisfaction, women scoring lower in both variables. Commitment level did not differ significantly between partners. Furthermore, Table 1 lists correlations between men and women; we found medium to large correlations for the study variables, confirming our first hypothesis. Commitment correlated significantly with relationship satisfaction and with dyadic coping, showing me-

							Correlations						
	Women		Men				Common dyadic coping		Commitment		Relationship satisfaction		
	M	SD	M	SD	t	d	actor effect	partner effect	actor effect	partner effect	actor effect	partner effect	
Common dyadic coping	3.83	.86	4.02	.77	3.16**	23		.46**	.39**	.18*	.53**	.45**	
Commitment	6.42	.69	6.32	.80	1.59	.14	.41**	.25**		.27**	.62**	.32**	
Relationship satisfaction	4.30	.53	4.41	.51	3.24**	21	.56**	.42**	.55**	.30**		.62**	

Table 1. Means, standard deviations, and correlations

Note. Range for common dyadic coping: 1-5, range for commitment: 1-7, range for relationship satisfaction 1-5. M = mean, SD = standard deviations, t = t-value, d = Cohen's d, N = 201 dyads; * p < .05, ** p < .01. Correlations between the same construct across partners are depicted on the main diagonal of the correlation table; correlations of women's variables with their own (actor) and their partner's (partner) scores are depicted below the main diagonal, correlations of men's scores with their own (actor) and their partner's (partner) scores above the main diagonal.

dium positive correlation coefficients for within-subject effects. Moreover, small to medium positive correlation coefficients were found for between-subject effects.

Multigroup Modeling

In order to investigate whether the age or relationship length are possible moderators influencing the model parameters in the CFM and APIM, we first ran the CFM and APIM using the total sample and in a second step in a multigroup approach. The first model to be run was the model with completely restricted structural parameters (only means and intercepts may differ between groups). For the CFM we found that the two completely restricted multigroup models showed an excellent fit to the data (two age cohorts: $\chi^2 = 48.57$; df = 45; p = .33; RMSEA = .03 [90%CI: 0 to .07]; CFI = .99; two relationship length groups: χ^2 = 59.79; df = 45; p = .07; RMSEA = .06 [90%CI: 0 to .09]; CFI = .97). The two completely restricted multigroup APIMS also showed an excellent fit to the data (two age cohorts: $\chi^2 = 44.02$; df = 42; p = .38; RMSEA = .02 [90%CI: 0 to .07]; CFI = .99; two relationship length groups: χ^2 = 56.92; df = 42; p = .06; RMSEA = .06 [90%CI: 0 to .10]; CFI = .97). Consequently, we report the model parameters using the total sample (1-group solutions).

The Common Fate Model

To test our second hypothesis of mediation on a dyadic level, we calculated the CFM. The model (see Figure 2) showed a good fit with $\chi^2(3) = .591$, p = .898, the comparative fit index (CFI) = 1.000 and root mean square error of approximation (RMSEA) < .001. Additionally, bootstrap analyses which were also used to test the indirect effect between the two latent variables revealed that the model showed an excellent fit to the data (Bollen-Stine bootstrap p-value of .908; if Bollen-Stine bootstrap p < .05, model is rejected, Bollen & Stine, 1992). Both effects from commitment to relationship satisfaction and from relationship satisfaction to common dyadic coping were significant. Moreover, the indirect effect proved to be significant, whereas the direct effect from commitment on common dyadic coping became nonsignificant. These results suggest that the association between intimate couples' relationship commitment and their common dyadic coping performance is fully mediated by their relationship satisfaction, which we stated in our hypothesis. Commitment explained 57% of the variance in relationship satisfaction, and both commitment, and relationship satisfaction accounted for a total of 70% of the variance in common dyadic coping.

The Actor Partner Mediation Model

For our third hypothesis, we calculated an APIMeM. In order to have the most parsimonious model, we used a stepwise modeling procedure. In a first step, we estimated the saturated model, that is, direct paths of exogenous variables to mediators and to endogenous variables as well as the mediation paths were included into the model. Within this model, none of the direct paths between commitment (exogenous variable) and dyadic coping (endogenous variable) was found to be significant. This led us to estimate three additional models: (1) a model without direct partner effects of commitment on dyadic coping, (2) a model without direct actor effects of commitment on dyadic coping, and (3) a model without any (actor or partner) effects of commitment on dyadic coping. The model without the direct partner effects (see Figure 3) showed the best fit, $\chi^2(2)$ = 2.002, p = .368; CFI = 1.000; RMSEA = < .002. To beable to fully interpret the patterns of this model and to test the mediating effects, we performed the Sobel test (1982) to test the indirect effects between all variables (see Ta-

Testing the Indirect Effects

Six of the eight indirect effects were significant (see Table 2); both men's and women's relationship satisfaction

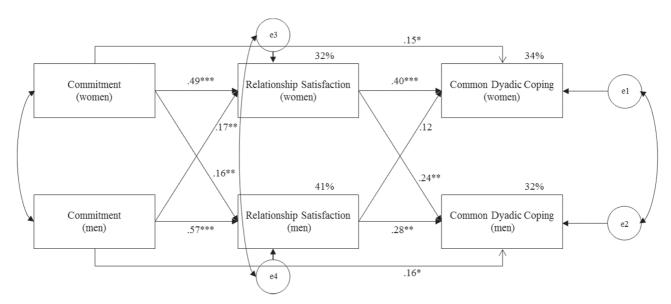


Figure 3. Actor-partner interdependence mediation model with commitment as endogenous variable, relationship satisfaction as mediator, and common dyadic coping as outcome variable. Note: e1 to e4 = residuals. Single-headed arrows = predictive paths, double-headed arrows = correlated variables. Coefficients represent standardized regression coefficients; percentages refer to amount of explained variances. ***p < .001, **p < .05.

Table 2. Mediated effects for actor-partner interdependence mediation model

Estimate	SE		
$a \times b$	$a \times b$	Z	р
.26	.06	4.22	.000
.14	.04	3.04	.002
.02	.02	1.44	.151
.05	.02	2.12	.034
.07	.03	2.40	.016
.04	.02	2.11	.035
.07	.05	1.64	.101
.15	.05	3.01	.003
	a × b .26 .14 .02 .05 .07 .04 .07	a × b a × b .26 .06 .14 .04 .02 .02 .05 .02 .07 .03 .04 .02 .07 .05	a × b a × b z .26 .06 4.22 .14 .04 3.04 .02 .02 1.44 .05 .02 2.12 .07 .03 2.40 .04 .02 2.11 .07 .05 1.64

Note. EX: exogenous variable; Med: Mediator; EN: endogenous variable; ComSec = cognitive and emotional commitment; RAS = relationship satisfaction; DC = dyadic coping; $(a \times b)$ = indirect effect consisting of 2 direct effects a and b; w = women; m = men; SE = standard error; z = z-scores; p = level of significance. Women's and men's commitment as independent variables, women's and men's relationship satisfaction as mediator variables, and women's and men's assessment of their common dyadic coping efforts as outcome variable.

mediates the association between their own commitment and their own dyadic coping (actor effects of mediation). Women's relationship satisfaction proved to significantly mediate between all indirect exogenous-endogenous paths in the model, which supports our third hypothesis. In other words, women's relationship satisfaction partially mediates between women's commitment and women's dyadic coping. Furthermore, it mediates between women's commitment and men's dyadic coping, between men's own commitment and women's dyadic coping —

and even between men's commitment and men's dyadic coping, although this latter mediational effect was weaker than the mediation through men's relationship satisfaction. Men's relationship satisfaction partially mediates between their own commitment and their own dyadic coping, and between women's commitment and men's dyadic coping, but not between women's commitment and women's dyadic coping and not between their own commitment and women's dyadic coping.

Discussion

Our study explored the association between long-term relationships' commitment, relationship satisfaction, and common dyadic coping in the realm of dyadic analyses, accounting for nonindependence of distinguishable dyads' data. Our data of 201 intimate couples, with a relationship duration of at least 10 years, revealed that women reported a significantly lower level of common dyadic coping than men, and that their relationship satisfaction was significantly lower than men's relationship satisfaction. Differences were significant, though we must note that the levels – particularly the relationship satisfaction scores – were high on average, indicating an overall satisfied study sample. Men and women did not differ significantly in their levels of commitment.

Before examining the mediation process, we conducted correlational analyses with our study variables, which revealed that higher levels of relationship commitment correlated significantly with both higher levels of relationship satisfaction and higher levels of common dyadic coping. Not only did we find significant actor effects between these variables, we also found significant partner effects between all variables. These results were a promising hint as to our hypothesized associations between the variables (Hypothesis 1) as well as a confirmation of nonindependence of dyadic data.

In a second step, we conducted an analysis using the common fate mediation model, looking at how strongly the correlations between both partners' manifest indicators could be attributed to the common dyadic latent variable, i.e., the common fate, and whether relationship satisfaction would mediate the association between the two latent constructs of commitment and dyadic coping. This model confirmed that relationship satisfaction fully mediated the association between intimate partners' cognitive and emotional commitment and their common dyadic coping. We found evidence in line with previous research findings, confirming the association between high levels of commitment and high relationship satisfaction (Acker & Davis, 1992; Lemieux & Hale, 1999). Our results support findings of relationship satisfaction's strong association with social support (Røsand, Slinning, Eberhard-Gran, Røysamb, & Tambs, 2012), and fostering the will to invest into the relationship (Wieselquist et al., 1999).

We then continued our analyses by using the APIMeM (Ledermann et al., 2011), which enabled us to thoroughly explore the role of relationship satisfaction as a hypothesized mediator between the exogenous variable of relationship commitment and the endogenous variable of assessed common dyadic coping on the level of the dyad members. Mediation analysis revealed classical actor-actor mediations: Women's relationship satisfaction partially mediated the association between their commitment and their common dyadic coping, and men's relationship satisfaction partially mediated the association between their commitment and their assessment of common dyadic coping. This latter association was mediated by women's relationship satisfaction, too. Based on findings by Davila et al. (1997), we assume that women's relationship satisfaction plays such a significant role in the mediating process between men's commitment and dyadic coping, because of the prominent effect that women's affect has on relationship functioning. Those authors found direct influence of women's dysphoria on their own as well as on their partners' social support behavior. In our study, direct, indirect and mediating effects of women's relationship satisfaction on dyadic coping behavior substantiate these findings.

Besides the significant insight the analysis of our data provided, we must discuss some limitations. An important caveat is that, because of the cross-sectional character of our data, our analyses cannot confirm a definite causality of our variables. Based on the considerations that led us to our model of the reinforcing role of commitment, we are convinced that commitment plays a major role in intimate relationships and – encouraged by our findings – especially in predicting intimate couples' prorelationship

behavior. Yet we must acknowledge that a group comparison with a clinical sample would contribute substantially to the relevance of our results. Because commitment has proven to also function as an endogenous variable, we would like to emphasize that, in the present study, we fully concentrated on one specific aspect of the potential processes between relationship variables of commitment, relationship satisfaction and dyadic coping. The direction of the hypothesized paths in our models is fully in line with the "cyclical growth" between dependence, commitment, and investments in long-term intimate relationships (Rusbult et al., 2001, p. 376).

Furthermore, we opted for very specific items of dyadic coping, rendering the true common aspect of items that are being assessed by both dyad members. We did this in order to take adequate account of the dyadic analysis level of the common fate model. However, to expand and consolidate our findings, further analyses should examine to which extent they apply to the items of supportive dyadic coping and even on negative dyadic coping on the level of dyad members or to extend the present research to the items of common dyadic coping concerning sexual/intimate behavior. It could then be of high interest to include the sexual commitment items of the Comsec, which we omitted in our study. Finally, it would be sensible to supplement our analyses with observer reports of specific affects in intimate couples' interactions (e.g., Johnson, 2002; Yoshimoto, Shapiro, O'Brien, & Gottman, 2005) in order to counter the risk of social desirability that the selfreports used in this study might be subject to.

Apart from these limitations, we believe that we have gained interesting insight into the complex mechanisms that connect variables of intimate relationships. The modeled structure between commitment and dyadic coping proposed in this study shows strong parallels to a dimension of commitment that Schoebi, Karney, and Bradbury (2012) defined, "as an inclination to engage in maintenance behaviors," and that was found to bear stabilizing influence on couples' relationship. Understanding the facets of commitment and their influence on couples' willingness to jointly cope with problems in order to maintain their relationships also has practical implications: In a clinical setting of couple counseling, therapists might use this knowledge to help clients identify and redefine their concepts of commitment, focusing on continuance and a common future as a couple (see Pope, 2013, for more de-

Finally, much of previous research has concentrated on individual analysis of commitment (Rusbult, Johnson, & Morrow, 1986; Slotter et al., 2012), relationship satisfaction and also on partners' individually assessed prorelationship efforts so far (Dainton & Aylor, 2002; Rabby, 2007). Therefore, this study contributes an important aspect by taking into account dyadic analysis, treating intimate couples as one entity, and looking in depth at non-independent processes on the level of dyad members.

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Declaration of Conflicts of Interest

The authors declare that there is no conflict of interest.

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